

## Qualifying Exam: 2025 Spring

考试课程: Probability & Statistics    姓名: \_\_\_\_\_    学号: \_\_\_\_\_

- There are 11 problems in this exam (3 pages). You need to choose 8 of them to solve. If you select more than 8, only the first 8 that you have worked on will be graded. Note that 4 of the problems are worth 15 points each and the rest 10 points each.
- You must follow all the rules of exam taking. Misconducts will be subject to proper disciplinary actions by the Center.
- You must provide all necessary details for full credits. A final answer with no or little explanation/derivation, even if correct, receives a minimal credit.

1. (10 points) Let  $\{U_n\}$  be i.i.d. with uniform distribution on  $[0, 1]$ . For  $n \geq 2$ , define

$$X_n := \min\{U_1, \dots, U_n\}, \quad Y_n = \max\{U_1, \dots, U_n\}.$$

- (1) Derive the joint distribution of  $(X_n, Y_n)$ .
- (2) Calculate  $\mathbb{E}[X_n | \sigma(Y_n)]$ .

2. (10 points) Suppose  $(X, Y, Z)$  is three-dimensional mean-zero Gaussian vector with covariance matrix given by

$$\begin{pmatrix} 1 & \rho & \tau \\ \rho & 1 & 0 \\ \tau & 0 & 1 \end{pmatrix}.$$

- (1) Show that  $\rho^2 + \tau^2 \leq 1$ .
  - (2) Show that there exist real numbers  $a, b, c$  and a standard Gaussian random variable  $W$  independent of  $(Y, Z)$  such that  $X = aY + bZ + cW$  and derive  $a, b, c$ .
3. (15 points) Suppose  $\{\xi_i\}$  are i.i.d. Bernoulli random variables with  $\mathbb{P}[\xi_1 = 1] = \mathbb{P}[\xi_1 = -1] = 1/2$ . Set  $S_n = \sum_{j=1}^n \xi_j$  and define  $\tau = \min\{n : S_n = -a \text{ or } b\}$  for  $a, b \in \mathbb{Z}_{>0}$ .

- (1) Calculate  $\mathbb{E}[e^{\lambda S_n}]$  for  $\lambda \in \mathbb{R}$ .
- (2) Calculate  $\mathbb{E}[s^\tau]$  for  $s \in (0, 1)$ .

4. (10 points) Suppose  $\{\xi_i\}$  are i.i.d. Bernoulli random variables with  $\mathbb{P}[\xi_1 = 1] = \mathbb{P}[\xi_1 = -1] = 1/2$ . Set

$$X_n = \sum_{j=1}^n 2^{j-1} \xi_j.$$

- (1) Define  $T = \min\{j \geq 1 : \xi_j = +1\}$ . Derive the law of  $T$  and calculate  $\mathbb{E}[T]$ .
- (2) Show that  $\{X_n\}$  is a martingale and calculate  $\mathbb{E}[X_T]$ .

5. (10 points) Let  $\{B_t\}_{t \geq 0}$  be a one-dimensional Brownian motion starting from the origin (i.e.,  $B_0 = 0$ ). For  $a \in (0, \infty)$ , define  $T_a := \inf\{t \geq 0 : |B_t| = a\}$ .

- (1) Prove that  $T_a < \infty$  a.s.

(2) Prove that  $T_a$  and  $\mathbf{1}_{\{B_{T_a}=a\}}$  are independent.

6. (15 points) Suppose  $X$  and  $Y$  are i.i.d with  $\mathbb{E}[X] = 0$  and  $\text{var}(X) = 1$ . Suppose further that

$$\mathbb{E}[(X - Y)^2 | \sigma(X + Y)] = 2.$$

Show that  $X$  and  $Y$  are independent standard Gaussian random variables.

7. (10 points) Let  $X_1, \dots, X_n$  ( $n \geq 2$ ) be iid uniform  $U(0, \theta)$  where  $0 < \theta < \infty$ .

(a) Under squared error loss  $l(\theta, \delta(\mathbf{X})) = (\delta(\mathbf{X}) - \theta)^2$ , derive the generalized Bayes rule for the improper prior

$$\pi(\theta) = 1, 0 < \theta < +\infty.$$

(b) Is the Bayes rule derived in (a) consistent for  $\theta$ ?

8. (10 points) Let  $X_1, \dots, X_n$  be iid Gamma(2025,  $\theta$ ) with  $\theta > 0$  unknown. Recall that the probability density function of a Gamma( $\alpha, \theta$ ) is  $f(x | \alpha, \theta) = \frac{1}{\Gamma(\alpha)\theta^\alpha} x^{\alpha-1} e^{-\frac{x}{\theta}}, x > 0$ , where  $\Gamma(\alpha) = (\alpha - 1)!$  for a positive integer  $\alpha$ .

(a) Consider the hypothesis testing problem  $H_0 : \theta = 1$  vs.  $H_1 : \theta \neq 1$ . Derive the UMP level  $\alpha = 0.05$  test if it exists. If it doesn't exist, please argue why it does not exist.

(b) Derive the MLE for  $\frac{1}{\theta}$ , what is the asymptotic distribution of this MLE for  $\frac{1}{\theta}$ ?

9. (15 points) Let  $(X_1, Y_1), \dots, (X_n, Y_n)$  be iid random bivariate vectors. Let  $\rho$  be the correlation coefficient between  $X_i$  and  $Y_i$ . Let

$$\hat{\rho} = \frac{1}{(n-1)\sqrt{S_X^2 S_Y^2}} \sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y}),$$

where  $\bar{X} = \frac{\sum_{i=1}^n X_i}{n}$ ,  $\bar{Y} = \frac{\sum_{i=1}^n Y_i}{n}$ ,  $S_X^2 = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n-1}$ ,  $S_Y^2 = \frac{\sum_{i=1}^n (Y_i - \bar{Y})^2}{n-1}$ .

(a) Assume that  $\mathbb{E}|X_i|^4 < \infty$  and  $\mathbb{E}|Y_i|^4 < \infty$ , prove that  $\sqrt{n}(\hat{\rho} - \rho) \rightarrow N(0, c)$  for some value  $c$ . Please identify  $c$  in terms of moments of  $(X_i, Y_i)$ , however, there is no need to simplify your calculation on  $c$ .

(b) Further, assume that  $(X_i, Y_i)$  follows a bivariate normal distribution that

$$\begin{pmatrix} X_i \\ Y_i \end{pmatrix} \sim N \left( \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix} \right).$$

The probability density function of the above bivariate normal distribution is given by

$$f(x, y | \rho) = \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left\{-\frac{1}{2(1-\rho^2)}[x^2 - 2\rho xy + y^2]\right\}.$$

Find a minimal sufficient statistic for  $\rho$ . Is this minimal sufficient statistic complete or not? Please explain.

10. (10 points) Let  $X_1, \dots, X_n$  be iid from  $N(0, \sigma^2)$  with  $\sigma > 0$  unknown. Let  $W = \frac{1}{n} \sum_{i=1}^n |X_i|$ . Recall  $\chi_k^2$  has probability density function

$$\frac{1}{2^{\frac{k}{2}} \Gamma(\frac{k}{2})} x^{\frac{k}{2}-1} e^{-\frac{x}{2}}, x \geq 0.$$

(a) Find an unbiased estimator of  $\sigma$  based on  $W$  and identify its limiting distribution.

- (b) Is it admissible under the square error loss? Justify your answer.
11. (15 points) Let  $X_1, \dots, X_n$  be iid from exponential distribution with density  $f(x; \lambda, \alpha) = \lambda e^{-\lambda(x-\alpha)} I_{\{x \geq \alpha\}}$ , where  $\lambda > 0$  and  $-\infty < \alpha < \infty$  are unknown parameters.
- (a) Find the MLE of  $(\lambda, \alpha)$ , say  $(\hat{\lambda}, \hat{\alpha})$ .
- (b) What distribution does  $\frac{\sqrt{n}(\hat{\lambda}-\lambda)}{n(\hat{\alpha}-\alpha)}$  converge to? You will receive partial credits if you can identify the distribution(s) of the numerator and/or denominator.